

A practical guide to computer simulation II

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July 3, 2003

8 Random Numbers

Examples for Random numbers used in computer simulations:

- Instances with quenched disorder, e.g. spin glasses (interactions are random)
- Simulation at finite temperatures using Monte Carlo algorithms
- Randomized algorithms (deterministic algorithms made random)

Literature: [1, 2].

8.1 Generating random numbers

Computers are deterministic \rightarrow no true randomness possible.

Randomness created by user (time intervals between keystrokes): not controllable.

Pseudo random numbers: generated deterministically but look random, i.e. have many properties of random numbers: uniform distribution, low correlations.

Linear congruential generator: generates sequence I_1, I_2, \dots between 0 and $m-1$, starting from given I_0 .

$$I_{n+1} = (aI_n + c) \bmod m \quad (1)$$

Random numbers r uniformly in interval $[0, 1)$: $r_n = I_n/m$. Arbitrary distributions (see below).

Task: choose parameters a, c, m (and I_0) such that generator is “good” \rightarrow test criteria needed. Attention: several times results for simulations were wrong because of bad random number generators [3].

Example: $a = 12351, c = 1, m = 2^{15}$ and $I_0 = 1000$ (and dividing by m) is “uniformly” distributed in $[0, 1)$ (Fig. 1).

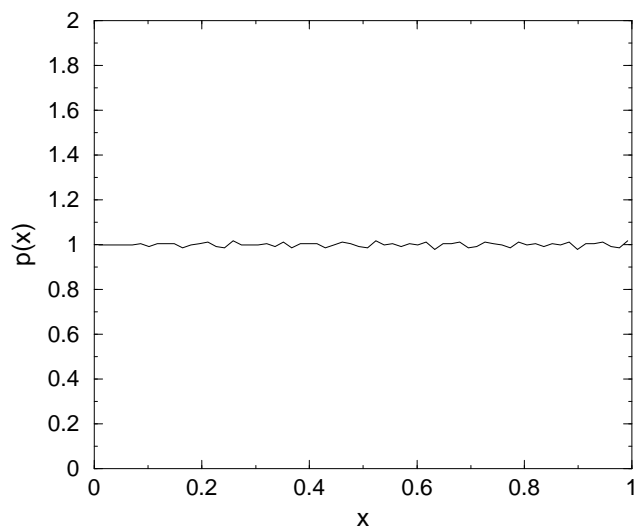


Figure 1: Distribution of random numbers in the interval $[0, 1)$. They are generated using a linear congruential generator with the parameters $a = 12351, c = 1, m = 2^{15}$.

But they have correlations. Study: k -tuples of k successive random numbers $(x_i, x_{i+1}, \dots, x_{i+k-1})$. Low correlations: k -dim space uniformly filled. LCGs: points lie on $(k - 1)$ -dimensional planes, number is *at most* $O(m^{1/k})$. Above case: few planes, see Fig. 2

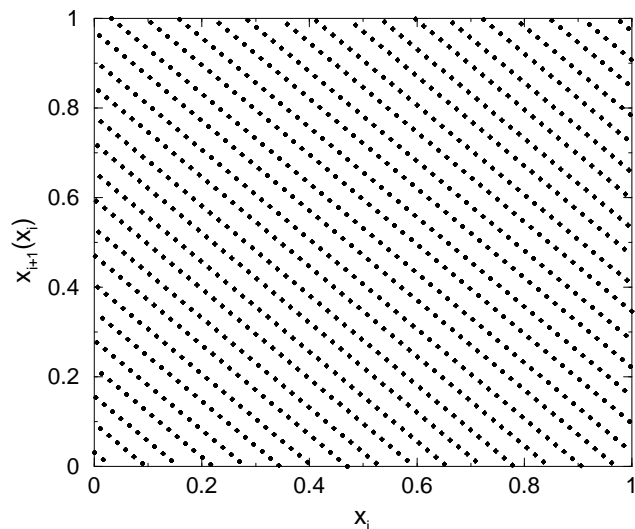


Figure 2: Two point correlations $x_{i+1}(x_i)$ between successive random numbers x_i, x_{i+1} . Linear congruential generator with the parameters $a = 12351, c = 1, m = 2^{15}$.

Better: $a = 12349$ instead, Fig. 3.

“Good generator”: $a = 7^5 = 16807, m = 2^{31} - 1, c = 0$. Note: more than 32-bit arithmetic needed, see Ref. [1].

Low-order bits much less random than the high-order bits \rightarrow numbers in an interval $[1, N]$:

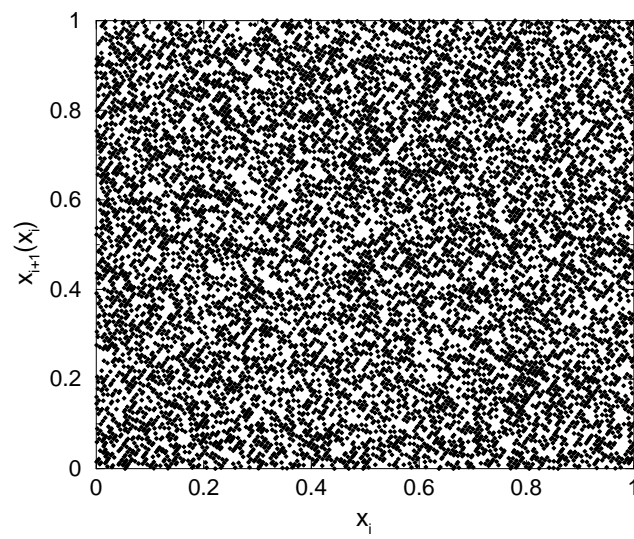


Figure 3: Two point correlations $x_{i+1}(x_i)$ between successive random numbers x_i, x_{i+1} . Linear congruential generator with the parameters $a = 12349, c = 1, m = 2^{15}$.

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r = 1+(int) (N*(I_n)/m);
```

instead of using the modulo

8.2 Inversion Method

Given: `drand()` generating uniformly distributed random numbers in $[0, 1)$.

Aim: random numbers Z according pdf $p(z)$ with distribution

$$P(z) \equiv \text{Prob}(Z \leq z) \equiv \int_{-\infty}^z dz' p(z') \quad (2)$$

target: find a function $g(X)$, such that after the transformation $Z = g(U)$. Assume g is strongly monotonically increasing i.e. can be inverted \rightarrow

$$P(z) = \text{Prob}(Z \leq z) = \text{Prob}(g(U) \leq z) = \text{Prob}(U \leq g^{-1}(z)) \quad (3)$$

Since $\text{Prob}(U \leq u) = F(u) = u$ for U uniformly in $[0, 1)$ and with identifying u with $g^{-1}(z)$, we get $P(z) = g^{-1}(z)$, hence $g(z) = P^{-1}(z)$. Works if P can be calculated (eventually numerically) and can be inverted.

Example:

Exponential distribution: probability density $p(z) = \lambda \exp(-\lambda z)$ with $P(z) = 1 - \exp(-\lambda z)$. Hence generate uniformly distributed random numbers U and choose $Z = -\ln(1 - U)/\lambda$.

8.3 Rejection Method

For non-integrable pdfs of non-invertible distribution functions, if distribution $p(z)$ fits into a box $[x_0, x_1] \times [0, p_{\max}]$, i.e. $p(z) = 0$ for $z \notin [x_0, x_1]$ and $p(z) \leq p_{\max}$. basic idea: generate random pairs (x, y) , which are distributed uniformly in $[x_0, x_1] \times [0, p_{\max}]$ and accept only those values x where $y \leq p(x)$ holds, i.e. the pairs which are located below $p(x)$, see Fig. 5.

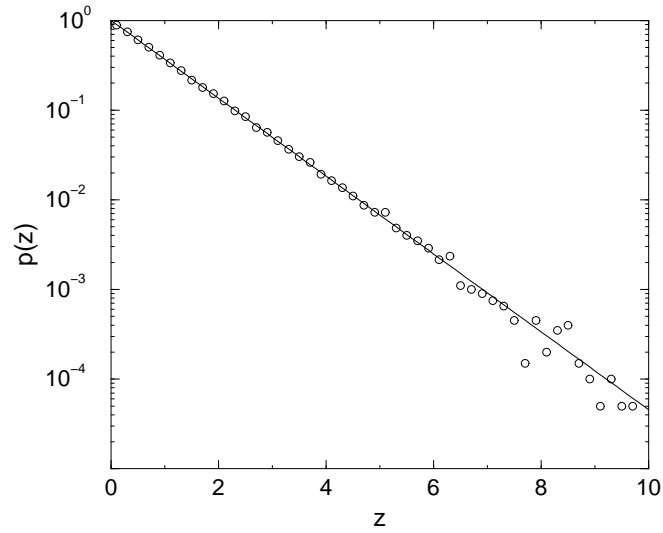


Figure 4: Histogram of random numbers generated according to an exponential distribution ($\lambda = 1$) compared with the probability density (straight line) in a logarithmic plot.

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algorithm rejection_method( $p_{\max}, x_0, x_1, p$ )
begin
   $found := \mathbf{false}$ ;
  while not  $found$  do
    begin
       $u_1 :=$  random number in  $[0, 1)$ ;
       $x := x_0 + (x_1 - x_0) \times u_1$ ;
       $u_2 :=$  random number in  $[0, 1)$ ;
       $y := p_{\max} \times u_2$ ;
      if  $y \leq p(x)$  then
         $found := \mathbf{true}$ ;
      end;
    return( $x$ );
  end

```

Drawback: Many random numbers needed, some are even thrown away.

8.4 The Gaussian Distribution

Gaussian distribution with mean m and width σ (most commonly used distribution in simulations), pdf: σ is (see also Fig. 6)

$$p_G(z) = \frac{1}{\sqrt{2\pi}\sigma} \exp\left(-\frac{(z-m)^2}{2\sigma^2}\right) \quad (4)$$

Here, z according normal distribution ($m = 0, \sigma = 1$). General case: use $\sigma z + m$
 Neither inversion nor rejection method works here. 3 possibilities

- Work artificially boxed, e.g. in $[-3, 3]$ \rightarrow no large values.

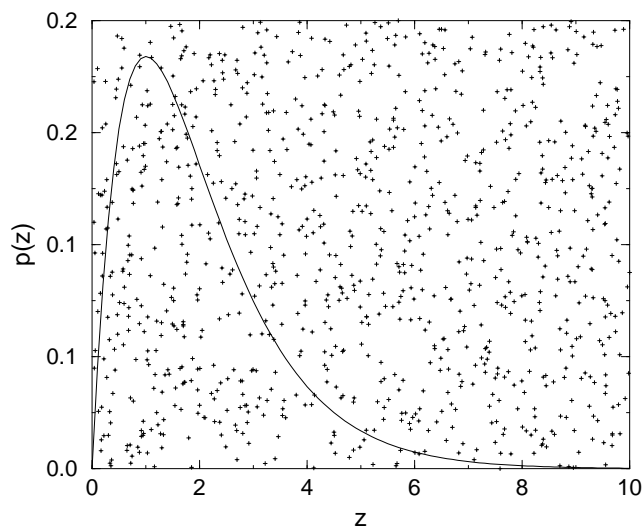


Figure 5: The rejection method: points (x, y) are scattered uniformly over a bounded rectangle. The probability that $y \leq p(x)$ is proportional to $p(x)$.

- Use central limit theorem: sum of N independently distributed random variables u_i (with mean m and variance v) converge to a Gaussian distribution with mean Nm and variance Nv . Use u_i ($m = 0.5, v = 1/12$) uniformly in $[0, 1)$, $N = 12$, then $Z = \sum_{i=1}^{12} u_i - 6$ as desired. Drawback: 12 numbers needed and boxed in $[-6, 6]$.
- *Box-Müller method*: take two values drawn from uniformly in $[0, 1)$ distributed random variables and set.

$$\begin{aligned} n_1 &= \sqrt{-2 \log(1 - u_1)} \cos(2\pi u_2) \\ n_2 &= \sqrt{-2 \log(1 - u_1)} \sin(2\pi u_2) \end{aligned}$$

Proof [1, 2]: Write n_1, n_2 in polar coordinates (r, θ) , i.e. $(r, \theta) = f(n_1, n_2)$, the inverse is:

$$\begin{aligned} n_1 &= r \cos(\theta) \\ n_2 &= r \sin(\theta) \end{aligned}$$

We want to obtain the pdfs for (r, θ) . For the general case $(W, Z) = f(X, Y)$ with $p_{X,Y}$ being the (joint) pdf of (X, Y) we have $p_{W,Z}(w, z) = p_{X,Y}(f^{-1}(w, z))|\mathbf{J}^{-1}|$ with $|\mathbf{J}^{-1}|$ being the Jacobi determinant of the inverse Transformation.

Using

$$|\mathbf{J}^{-1}| = \begin{vmatrix} \frac{\partial n_1}{\partial r} & \frac{\partial n_1}{\partial \theta} \\ \frac{\partial n_2}{\partial r} & \frac{\partial n_2}{\partial \theta} \end{vmatrix} = \begin{vmatrix} \cos(\theta) & -r \sin(\theta) \\ \sin(\theta) & r \cos(\theta) \end{vmatrix} = r \cos^2(\theta) + r \sin^2(\theta) = r \quad (5)$$

we get

$$p_{R,\Theta}(r, \theta) = \frac{r}{2\pi} e^{-n_1^2/2 - n_2^2/2} = \frac{r}{2\pi} e^{-r^2/2} \quad (6)$$

The distribution factorizes in r and θ . Hence θ can be taken uniformly distributed in $[0, 2\pi]$ (i.e. $\theta = 2\pi u_2$) and $p_R(r) = r e^{-r^2/2}$ (*). Now it remains

to see how to generate random numbers according p_R .

For this, consider X exponentially distributed with parameter λ , i.e. $p_X(x) = \lambda e^{-\lambda x}$. Let's take $Y = \sqrt{X}$. We want to obtain the pdf $p_Y(y)$. For the general case $Y = H(X)$ (H strongly monotonic) it is $p_Y(y) = p_X(H^{-1}(y)) \frac{1}{|H'(H^{-1}(y))|}$. Here with $H(x) = \sqrt{x}$, i.e. $H^{-1}(y) = y^2$ and $H'(x) = 1/2\sqrt{x}$, we get $p_Y(y) = 2\lambda y e^{-\lambda y^2}$. Comparing with (*), we see that taking X exponentially distributed with $\lambda = 0.5$ (i.e. $x = -2 \log(1 - u_1)$) and then taking $r = \sqrt{x}$ we get the desired distribution for r . QED.

- Simulation of particles in a box is explained in Ref. [4]

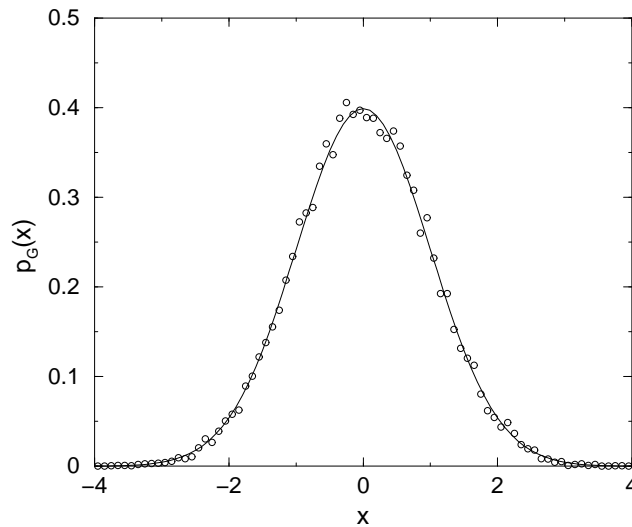


Figure 6: Gaussian distribution with zero mean and unit width. The circles represent a histogram obtained from 10^4 values drawn with the Box-Müller method.

8.5 Exercise

Write a program that generates 10^4 random numbers for a Gaussian distribution boxed in $[-6, 6]$ using the *rejection* method and record a normalized histogram of bin width 0.1. Draw the histogram and the Gaussian distribution in gnuplot.

References

- [1] W.H. Press, S.A. Teukolsky, W.T. Vetterling, and B.P. Flannery, *Numerical Recipes in C* (Cambridge University Press, Cambridge 1995)
- [2] B.J.T. Morgan, *Elements of Simulation*, (Cambridge University Press, Cambridge 1984)
- [3] A.M. Ferrenberg, D.P. Landau and Y.J. Wong, *Phys. Rev. Lett.* **69**, 3382 (1992); I. Vattulainen, T. Ala-Nissila and K. Kankaala, *Phys. Rev. Lett.* **73**, 2513 (1994)
- [4] J.F. Fernandez and C. Criado, *Phys. Rev. E* **60**, 3361 (1999)